

Defn: Expectation / Expected Value / Mean

(i) Discrete :  $E[X] = \sum_x x f(x)$

(ii) Continuous:  $E[X] = \int_R x f(x) dx$

Ex.  $X \sim \text{Bin}(n, p)$  independent

= flip  $n$  coins w/ prob  $p$  of H,  
 $X = \#$  of heads

$$\text{Support}(X) = \{0, 1, 2, \dots, n\}$$

the pmf is

$$f(x) = P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}$$

Note: To show this is a valid pmf

(1)  $f(x) \geq 0$

(2)  $\sum_x f(x) = 1$  (Binomial Theorem)

$$(x+y)^n = \sum_{x=0}^n (\dots)$$

$$E[X] = \sum_{x=0}^n x f(x)$$

$$= \sum_{x=0}^n x \binom{n}{x} p^x (1-p)^{n-x}$$

$$\underline{\underline{n}} \quad (n-1) \quad x, \quad \underline{\underline{n-x}}$$

aside:

$$\binom{n}{x} = \frac{n!}{x!(n-x)!}$$

$$\underline{\underline{n!}}$$

$$= \frac{n!}{(x-1)!(n-x)!}$$

$$x=0$$

$$= n \frac{(n-1)!}{(x-1)!(n-1-(x-1))!}$$

$$\text{define } y = x-1 \Leftrightarrow x = y+1$$

$$= \sum_{y=0}^{n-1} n \binom{n-1}{y} p^y (1-p)^{n-y-1} = n \binom{n-1}{x-1}$$

$$= np \sum_{y=0}^{n-1} \binom{n-1}{y} p^y (1-p)^{(n-1)-y}$$

pmf of  $\text{Bin}(n-1, p)$

$$Y \sim \text{Bin}(n-1, p)$$

$$f(y) = \binom{n-1}{y} p^y (1-p)^{(n-1)-y}$$

$$\text{Support}(Y) = \{0, 1, \dots, n-1\}$$

Sum of pmf of  $\text{Bin}(n-1, p)$   
over support

$$= 1$$

$$\boxed{= np} = (\# \text{ of trials}) (\text{prob. of success for each})$$

$\wedge$  doesn't have to be in the support  
(integer)

General trick:

Convert tricky sums/integrals to  
sums/integrals of pmf/pdfs.

Theorem: Law of the Unconscious Statistician

$$\mathbb{E}[g(X)] = \begin{cases} \sum_x g(x) f(x) & \text{discrete} \end{cases}$$

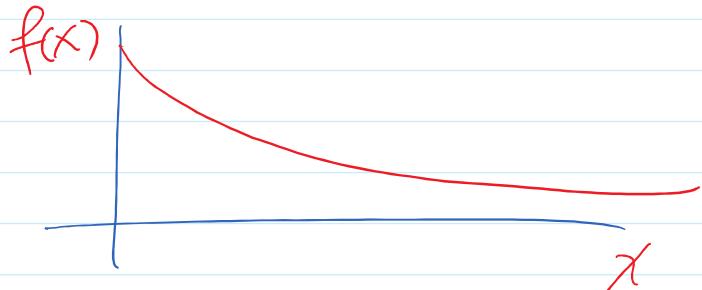
$$\int_R g(x) f(x) dx$$

Ex.  $X \sim \text{Exp}(\lambda)$

$$f(x) = \lambda e^{-\lambda x} \text{ for } x > 0$$

Saw

$$E[X] = \frac{1}{\lambda}$$



$$Q: E[X^2] = \int x^2 f(x) dx$$

$$= \int_0^\infty x^2 \lambda e^{-\lambda x} dx = \int u du$$

$$\text{by Parts: } u = x^2 \quad v = -e^{-\lambda x} \\ du = 2x dx \quad dv = \lambda e^{-\lambda x} dx$$

$$\begin{aligned} &= uv - \int v du \\ &= \left[ -x^2 e^{-\lambda x} \right]_0^\infty - \int_0^\infty -e^{-\lambda x} 2x dx \\ &= 0 + 2/\lambda \int_0^\infty x e^{-\lambda x} dx \end{aligned}$$

$$\text{d} \int E[X] = \int x f(x) dx$$

$$= \frac{2}{\lambda} \cdot \frac{1}{\lambda}$$

$$= \frac{2}{\lambda^2}$$

Q: Does expected value always exist? No.

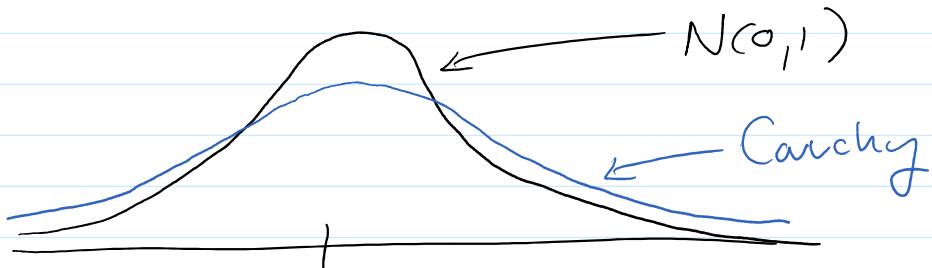
improper integral:

$$\int_0^\infty \frac{1}{x^2} dx \text{ has a meaningful value}$$

$$\int_0^\infty \frac{1}{x} dx \text{ no meaningful value}$$

Ex.  $X$  has a Cauchy distribution

$$f(x) = \frac{1}{\pi} \frac{1}{1+x^2} \text{ for } x \in \mathbb{R}$$



$$E[X] = \int x \frac{1}{1+x^2} dx \text{ doesn't converge} = \infty$$

$$E[X] = \int_{-\infty}^{\infty} x \frac{1}{1+x^2} dx \text{ doesn't converge} = \infty$$

### Theorem: Properties of E

①  $E[aX + b] = aE[X] + b$   $a, b \in \mathbb{R}$   
(linearity)

Pf.

$$\begin{aligned}
 E[aX + b] &= \int (ax + b) f(x) dx \\
 &= a \int x f(x) dx + \int b f(x) dx \\
 &= aE[X] + b \underbrace{\int f(x) dx}_I
 \end{aligned}$$

② If  $X \geq 0$  (Support(X) is non-neg.)  
then  $E[X] \geq 0$ .

Pf.  $E[X] = \int_0^\infty x f(x) dx \geq 0$

(c) If  $g_1$  and  $g_2$  are functions and  

$$g_1(x) \leq g_2(x) \quad \forall x$$

then

$$\mathbb{E}[g_1(X)] = \mathbb{E}[g_2(X)]$$

Pf. Combine (a) and (b)

$$\mathbb{E}[g_1(X)] \leq \mathbb{E}[g_2(X)]$$

$\Updownarrow$

$$\mathbb{E}[g_1(X)] - \mathbb{E}[g_2(X)] \leq 0$$

$\Updownarrow$  (linearity)

$$\mathbb{E}[\underbrace{g_1(X) - g_2(X)}_{\text{call this } Y}] \leq 0$$

$$\mathbb{E}[Y] \leq 0 \quad \text{where } Y \leq 0$$

(d) Furthermore if  $a \leq X \leq b$  then

$$a \leq \mathbb{E}[X] \leq b.$$

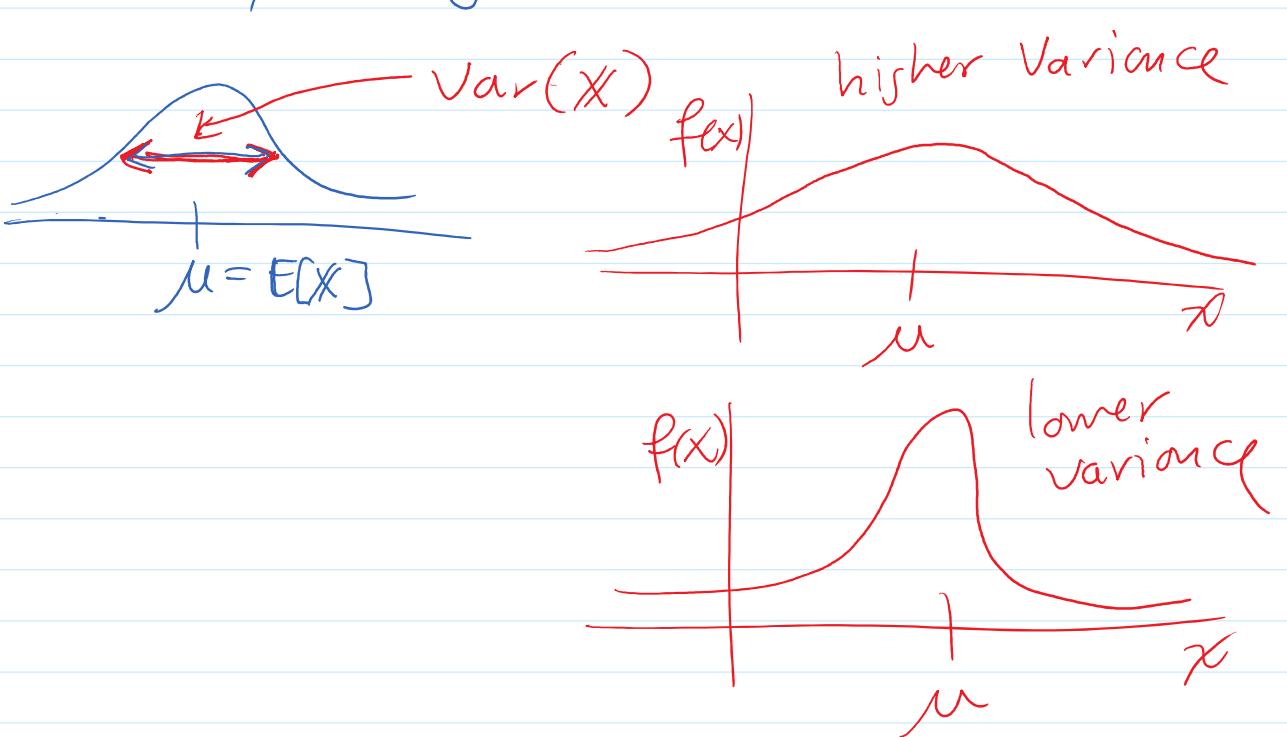
Pf use previous theorems.

Defn : Variance

Variance of a r.v. tells you how spread the mass/density is around the mean.

$$\sim \overbrace{\text{Var}(x)}$$

Higher Variance



Mathematically,

$$\boxed{\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2]}$$

Note:

$$\begin{aligned}
 Y &= X - \mathbb{E}[X] \\
 X - \mu &\quad \text{some number} \\
 \mathbb{E}[Y] &= \mathbb{E}[X - \mathbb{E}[X]] = \mathbb{E}[X] - \mathbb{E}[\mathbb{E}[X]] \\
 &= \mathbb{E}[X] - \mathbb{E}[X] \\
 &= 0
 \end{aligned}$$

Ex.  $X \sim \text{Exp}(\lambda)$  here  $\lambda > 0$

recall:  $\mathbb{E}[X] = \frac{1}{\lambda} = \mu$

Recall:  $E[X] = \lambda - \mu$

$$\begin{aligned}
 \text{Var}(X) &= E[(X-\mu)^2] \\
 &= \int (x-\mu)^2 f(x) dx \\
 &= \int (x-\frac{1}{\lambda})^2 \lambda e^{-\lambda x} dx \\
 &= \int \left( x^2 - \frac{2x}{\lambda} + \frac{1}{\lambda^2} \right) \lambda e^{-\lambda x} dx \\
 &= \underbrace{-\int x^2 \lambda e^{-\lambda x} dx}_{E[X^2]} - \frac{2}{\lambda} \underbrace{\int x \lambda e^{-\lambda x} dx}_{E[X]} + \frac{1}{\lambda^2} \underbrace{\int \lambda e^{-\lambda x} dx}_1 \\
 &= \frac{2}{\lambda^2} - \frac{2}{\lambda} \frac{1}{\lambda} + \frac{1}{\lambda^2} \\
 &= \frac{1}{\lambda^2} = \text{Var}(X).
 \end{aligned}$$

### Theorem 1: Short-cut Theorem For Variance

$$\text{Var}(X) = E[X^2] - [E[X]]^2$$

Pf.

$$\text{Var}(X) = E[(X-\mu)^2]$$

$$= E[X^2 - 2X\mu + \mu^2]$$

$E[c] = c$   
when  $c \in \mathbb{R}$

$$\begin{aligned}
 &= E[X^2] - E[2X\mu] + E[\mu^2] \quad \text{a const.} \\
 &= E[X^2] - 2\mu \underbrace{E[X]}_{\mu} + \mu^2 \\
 &= E[X^2] - 2\mu^2 + \mu^2 \\
 &\in E[(X^2) - \mu^2] \\
 &- E[X^2] - [E[X]]^2
 \end{aligned}$$

Ex.  $X \sim \text{Exp}(\lambda)$

$$E[X^2] = \frac{2}{\lambda^2} \quad \text{and} \quad E[X] = \frac{1}{\lambda}$$

so

$$\text{Var}(X) = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

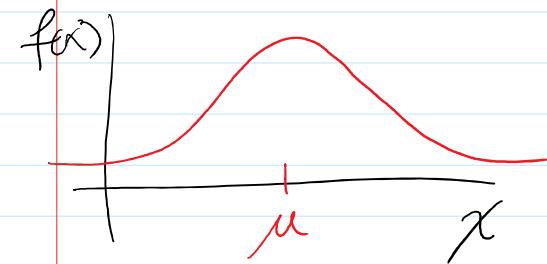
Theorem:

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

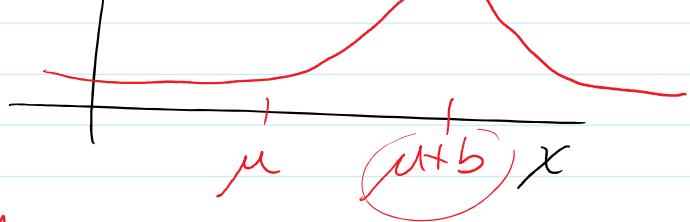
Pf.

$$\begin{aligned}
 &\text{Var}(aX + b) \\
 &= E[(aX + b)^2] - E[aX + b]^2 \quad (\text{short-cut}) \\
 &= E[a^2 X^2 + 2abX + b^2] - (a E[X] + b)^2 \\
 &= a^2 E[X^2] + 2ab E[X] + b^2 - (a^2 E[X]^2 + 2ab E[X] + b^2)
 \end{aligned}$$

$$\begin{aligned}
 &= a^2 E[X^2] - a^2 E[X]^2 \\
 &= a^2 (\underbrace{E[X^2] - E[X]^2}_{\text{(short-cut)}}) \quad (\text{short-cut}) \\
 &= a^2 \text{Var}(X)
 \end{aligned}$$



add  $b$



multiply  
a

