

pf of (3)

$$E[S^2] = E\left[\frac{1}{N-1} \sum_{n=1}^N (X_n - \bar{X})^2\right]$$

trick: $\text{Var}(Z) = E[Z^2] - E[Z]^2$

turns out:

$$\begin{aligned} \sum_n (X_n - \bar{X})^2 &= \sum_n X_n^2 - \left(\sum_n X_n\right)^2 \\ &= \sum_n X_n^2 - N(\bar{X})^2 \end{aligned}$$

$$= \frac{1}{N-1} E\left[\sum_n X_n^2 - N\bar{X}^2\right]$$

$$= \frac{1}{N-1} \left(\sum_n E[X_n^2] - N E[\bar{X}^2] \right)$$

re-arrange short-cut: $E[Z^2] = \text{Var}(Z) + E[Z]^2$

$$E[X_n^2] = \text{Var}(X_n) + E[X_n]^2$$

$$E[X_n^2] = \text{Var}(X_n) + E[X_n]^2 \\ = \sigma^2 + \mu^2$$

$$E[\bar{X}^2] = \text{Var}(\bar{X}) + E[\bar{X}]^2 \\ = \frac{\sigma^2}{N} + \mu^2$$

$$= \frac{1}{N-1} \left(\sum_n (\sigma^2 + \mu^2) - N \left(\frac{\sigma^2}{N} + \mu^2 \right) \right)$$

$$= \frac{1}{N-1} \left(N \cancel{(\sigma^2 + \mu^2)} - N \left(\frac{\sigma^2}{N} + \cancel{\mu^2} \right) \right)$$

$$= \frac{1}{N-1} (N-1) \sigma^2 = \sigma^2$$

Theorem: If $X_n \stackrel{iid}{\sim} f$ w/ MGF M

then

$$\underline{M_{\bar{X}}(t) = M(t/N)^N}$$

pf. Recall: $M(t) = E[e^{tX}]$

Facts: $M_{aX+b}(t) = e^{tb} M_X(at)$

If $A \perp B$ then $E[AB] = E[A]E[B]$

$$M_{\bar{X}}(t) = E[e^{t\bar{X}}] = E\left[e^{t \frac{1}{n} \sum X_n}\right]$$

$$= E\left[\prod_n e^{tX_n/n}\right]$$

$$= \prod_n \underbrace{E[e^{tX_n/n}]}_{M(t/n)}$$

$$= \prod_n M(t/n)$$

$$= M(t/n)^N$$

$$e^a e^b = e^{a+b}$$

$$\prod_i e^{a_i} = e^{\sum a_i}$$

$$= M(t/N)^N$$

Ex. $X_n \stackrel{iid}{\sim} \text{Gamma}(\alpha, \beta)$

What's the dist
of \bar{X} ?

$$f(x) = \frac{\beta^\alpha x^{\alpha-1} e^{-\beta x} \mathbb{1}(x > 0)}{\Gamma(\alpha)}$$

$$M(t) = (1 - t/\beta)^{-\alpha}$$

$$M_{\bar{X}}(t) = M(t/N)^N$$

$$= \left[(1 - t/N\beta)^{-\alpha} \right]^N$$

$$= (1 - t/N\beta)^{-\alpha N}$$

looks like
Gamma but
replace
 $\beta \rightarrow N\beta$
 $\alpha \rightarrow N\alpha$

So MGF of \bar{X} is the MGF of

$\text{Gamma}(N\alpha, N\beta)$

Gamma($N\alpha, N\beta$)

So \bar{X} has this dist.

Theorem: \bar{X} and S^2 for Normal Data

Let $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

✓ ① $\bar{X} \sim N(\mu, \sigma^2/N)$ (prove)

✓ ② $\bar{X} \perp S^2$ (later)

✓ ③ $\frac{N-1}{\sigma^2} S^2 \sim \chi^2(N-1)$ (sketch)

↑ chi-squared dist w/
 $N-1$ degrees of freedom

pp of ①

Use MGF theorem.

(MGF of $N(\mu, \sigma^2)$)

$$M(t) = \exp(\mu t + \frac{\sigma^2 t^2}{2})$$

Use MGF theorem. $M(t) = \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right)$

$$M_{\bar{X}}(t)$$

$$= M(t/N)^N$$

$$= \exp\left(\mu \frac{t}{N} + \frac{\sigma^2 (t/N)^2}{2}\right)^N$$

$$\left[\begin{aligned} (e^a)^b \\ = e^{ab} \end{aligned} \right]$$

$$= \exp\left(\mu t + \frac{\sigma^2 t^2}{2N}\right)$$

$$= \exp\left(\mu t + \frac{(\sigma^2/N)t^2}{2}\right)$$

↑ MGF of $N(\mu, \sigma^2/N)$

So $\bar{X} \sim N(\mu, \sigma^2/N)$

Chi-Squared dist

$\chi^2(k)$

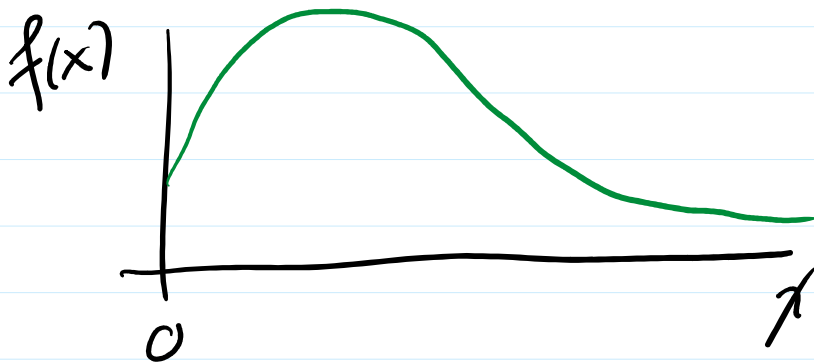
degrees of freedom
(k)

$$z \sim \chi^2(k)$$

degrees of freedom (dot)

$$k > 0$$

$$f(x) = \frac{1}{2^{k/2} \Gamma(k/2)} x^{k/2 - 1} e^{-x/2} \mathbb{1}(x > 0)$$



$$E[z] = k, \text{ Var}(z) = 2k$$

Facts:

① $z \sim N(0, 1)$ then $z^2 \sim \chi^2(1)$

② If $z_1 \sim N(0, 1), z_2 \sim N(0, 1), z_1 \perp z_2$

then $z_1^2 + z_2^2 \sim \chi^2(2)$

then $z_1^2 + z_2^2 \sim \chi^2(2)$

Generally, if $z_i \stackrel{iid}{\sim} N(0,1)$ then

$$\sum_{i=1}^N z_i^2 \sim \chi^2(N)$$

③ If $Y_n \stackrel{indep}{\sim} \chi^2(k_n)$ then

$$\sum_{n=1}^N Y_n \sim \chi^2\left(\sum_n k_n\right)$$

#3 in above theorem:

$$\frac{N-1}{\sigma^2} S^2 \sim \chi^2(N-1)$$

$$S^2 = \frac{1}{N-1} \sum_n (X_n - \bar{X})^2$$

$$\text{so } \frac{N-1}{\sigma^2} S^2 = \sum_n \frac{(X_n - \bar{X})^2}{\sigma^2}$$

kinda like

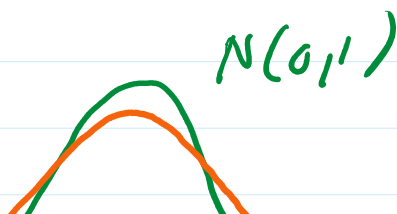
$$\begin{aligned}
 \text{So } \frac{N-1}{\sigma^2} S^2 &= \sum_n \frac{(x_n - \bar{x})^2}{\sigma^2} \\
 &= \sum_n \left(\frac{x_n - \bar{x}}{\sigma} \right)^2 \quad \text{Kinda like } \mu \\
 &\stackrel{\text{cheating}}{\approx} \sum_n \underbrace{\left(\frac{x_n - \mu}{\sigma} \right)^2}_{N(0,1)^2} \sim \chi^2(N)
 \end{aligned}$$

Penalty for replacing \bar{x} w/ μ is we lose a dof.

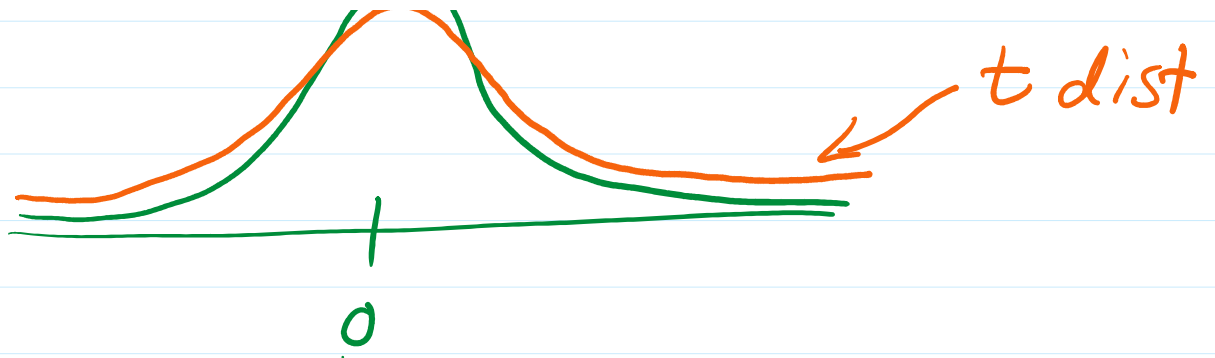
$$\text{So } \boxed{\frac{N-1}{\sigma^2} S^2 \sim \chi^2(N-1)}$$

t-distribution

like a $N(0,1)$ but w/ fatter tails



- t dist



One parameter k (called dof)

$$f(x) = \frac{\Gamma(\frac{k+1}{2})}{\sqrt{k\pi} \left(1 + \frac{x^2}{k}\right)^{\frac{k+1}{2}}}$$

$$\forall x \in \mathbb{R}$$

Fact: $U \sim N(0,1)$
 $V \sim \chi^2(k)$ $\Rightarrow U \perp V$

then $\frac{U}{\sqrt{V/k}} \sim \underline{t(k)}$

If $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

If $X_n \sim N(\mu, \sigma)$

① $\bar{X} \sim N(\mu, \sigma^2/N)$

② $\bar{X} \perp S^2$

③ $\frac{N-1}{\sigma^2} S^2 \sim \chi^2(N-1)$ $S = \sqrt{S^2}$

So $\frac{\bar{X} - \mu}{S/\sqrt{N}} \sim t(N-1)$

$Z \sim N(a, b^2)$

$\frac{Z - a}{b} \sim N(0, 1)$

pf $U = \frac{\bar{X} - \mu}{\sigma/\sqrt{N}} \sim N(0, 1)$

indep

$V = \frac{N-1}{\sigma^2} S^2 \sim \chi^2(N-1)$

$\frac{U}{\sqrt{V/k}} = \frac{\bar{X} - \mu}{\sigma/\sqrt{N}} \sqrt{\frac{N-1}{\sigma^2 S^2}} \sqrt{\frac{N-1}{N-1}}$

$$\Rightarrow \frac{\frac{\bar{X} - \mu}{\sigma/\sqrt{N}}}{\frac{\sqrt{S^2}}{\sigma}} = \frac{\bar{X} - \mu}{(S/\sqrt{N})} \sim t(N-1)$$

Generally, we'll work w/ parameterized families of dists, e.g.

- $N(\mu, \sigma^2)$, $\mu \in \mathbb{R}$, $\sigma^2 > 0$
- $\text{Exp}(\lambda)$, $\lambda > 0$
- $U(0, \theta)$, $\theta > 0$

Exponential Families

Assume we have a fam of dists parameterized by $\theta \in \Theta \subset \mathbb{R}$ so that $X_n \stackrel{\text{iid}}{\sim} f_\theta$

$$X_n \simeq F_\theta$$

where

$$f_\theta(\underline{x}) = h(\underline{x})c(\theta)\exp(T(\underline{x})w(\theta))$$

then we say the X_n 's are from
an exponential family of distr

ex. Poisson, Exp, Normal, Gamma, ...
