

Maximum Likelihood Estimation (MLE)

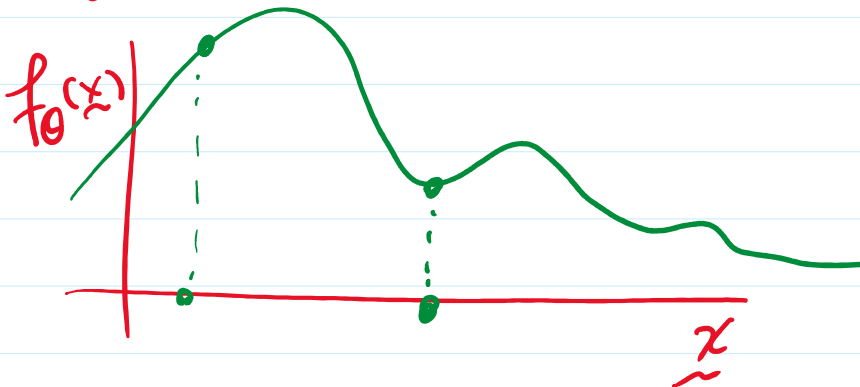
Assume $X_n \stackrel{iid}{\sim} f_\theta$ where $\theta \in \Theta$

recall: joint

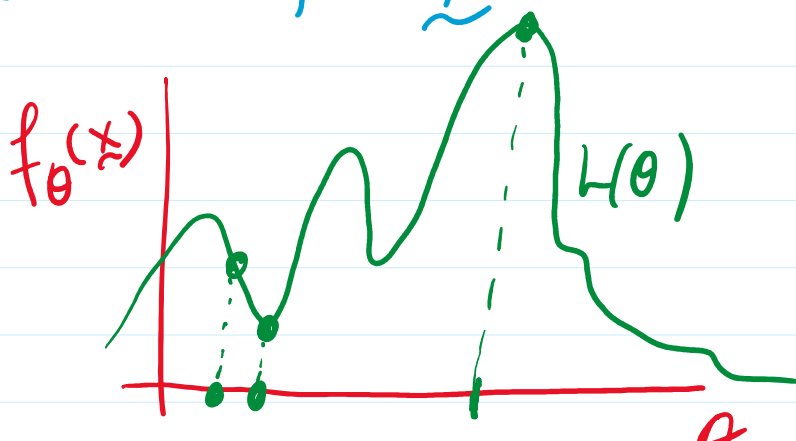
$$f_\theta(\underline{x}) = \prod_{n=1}^N f_\theta(x_n)$$

typ. look at as a function of \underline{x}

$$f_\theta: \mathbb{R}^N \rightarrow \mathbb{R}$$



Can also fix \underline{x} and think of as a fn of θ





Called the likelihood function

$$L: \Theta \rightarrow \mathbb{R}$$

defn:

$$L(\theta) = f_{\theta}(\underline{x})$$

MLE says choose $\hat{\theta}_{MLE}$ as val.
of θ that max. L .

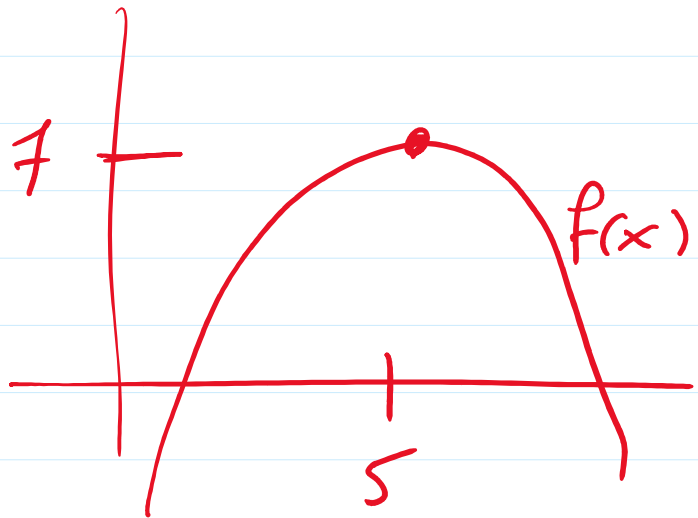
MLE:

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} L(\theta)$$

argmax

$$\max_x f(x) = 7$$

$$\operatorname{argmax}_x f(x) = 5$$



Often, we work with log-likelihood

$$l(\theta) = \log L(\theta).$$

Alt. defn of MLE

$$\hat{\theta}_{MLE} = \operatorname{argmax}_{\theta \in \Theta} l(\theta)$$

This is equiv. b/c log is inc. fn.

Ex, $X_n \stackrel{iid}{\sim} N(\theta, 1)$ where $\theta \in \mathbb{R}$.

... ..

What is the MLE?

① let's get $l(\theta)$

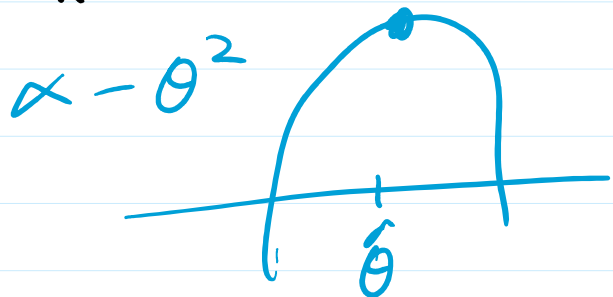
$$L(\theta) = f_{\theta}(\underline{x}) = \prod_{n=1}^N f(x_n)$$

$$= \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(x_n - \theta)^2\right)$$

$$= (2\pi)^{-N/2} \exp\left(-\frac{1}{2} \sum_{n=1}^N (x_n - \theta)^2\right)$$

$$l(\theta) = -\frac{N}{2} \log(2\pi) - \frac{1}{2} \sum_{n=1}^N (x_n - \theta)^2$$

② let's get $\frac{\partial l}{\partial \theta}$



$$\frac{\partial l}{\partial \theta} = 0 - \frac{1}{2} \sum_{n=1}^N 2(x_n - \theta)(-1)$$

$$= \sum_{n=1}^N (x_n - \theta)$$

$$= \sum_{n=1}^N (x_n - \theta)$$

3) set $\frac{\partial \ell}{\partial \theta} = 0$ and solve for θ

$$\frac{\partial \ell}{\partial \theta} = \sum_{n=1}^N x_n - N\theta = 0$$

$$\text{So } \hat{\theta}_{MLE} = \frac{1}{N} \sum_{n=1}^N x_n = \bar{x}$$

Technically, need to check

$$\frac{\partial^2 \ell}{\partial \theta^2} < 0 \quad \text{and consider } \lim_{\theta \rightarrow \pm\infty} L(\theta) = 0$$

Theorem: MLEs are functions of SS

$$\hat{\theta}_{MLE} = \text{function}(T)$$

↑
T suff.

- MLE

↑ T suff.
for θ .

pf: Factorization theorem says

$$L(\theta) = f_{\theta}(\underline{x}) = g(T, \theta) h(\underline{x})$$

So

$$\hat{\theta}_{MLE} = \operatorname{argmax}_{\theta \in \Theta} L(\theta)$$

$$= \operatorname{argmax}_{\theta \in \Theta} g(T, \theta) h(\underline{x})$$

$$= \operatorname{argmax}_{\theta \in \Theta} g(T, \theta)$$

↑ fn of T

$$\ell(\theta) = \log h(\underline{x}) + \log g(T, \theta)$$

$$\frac{\partial \ell}{\partial \theta} = \frac{\partial}{\partial \theta} \log g(T, \theta)$$

Q. $X_n \stackrel{iid}{\sim} \text{Bern}(p)$, $p \in [0, 1]$

What's my MLE for p ?

① Get $L(p)$ and $l(p)$

$$L(p) = f_p(\underline{x}) = \prod_{n=1}^N f(x_n)$$

$$= \prod_{n=1}^N p^{x_n} (1-p)^{1-x_n} \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$\bar{X} = \frac{1}{N} \sum_n x_n$$

$$N\bar{X} = \sum_n x_n$$

$$= p^{\sum_n x_n} (1-p)^{\sum_n (1-x_n)} \prod_n \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$= p^{N\bar{X}} (1-p)^{N(1-\bar{X})} \prod_n \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$l(p) = \log L(p)$$

$$= N\bar{X} \cdot \log(p) + N(1-\bar{X}) \log(1-p)$$

$$+ \log \prod_n \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$+ \log \prod_n \mathbb{1}(X_n = 0 \text{ or } 1)$$

② Take deriv, set eq to zero.

$$\frac{\partial \mathcal{L}}{\partial p} = \frac{N\bar{X}}{p} - \frac{N(1-\bar{X})}{1-p} = 0$$

$$\Rightarrow \frac{N\bar{X}}{p} = \frac{N(1-\bar{X})}{1-p}$$

$$\Rightarrow N\bar{X}(1-p) = N(1-\bar{X})p$$

$$\Rightarrow N\bar{X} - \cancel{N\bar{X}p} = Np - \cancel{N\bar{X}p}$$

$$\Rightarrow N\bar{X} = Np$$

$$\Rightarrow \boxed{\hat{p}_{MLE} = \bar{X}} = \text{pct of data that are 1s.}$$

— 1-1 for w/d

Consider

$$\eta = \frac{p}{1-p} = \text{odds}$$

$$p = \frac{\eta}{(1+\eta)}$$

Might want $\hat{\eta}_{MLE}$.

Get likelihood in terms of η ...

$$L(p) = p^{N\bar{x}} (1-p)^{N(1-\bar{x})}$$

$$L(\eta) = \left(\frac{\eta}{1+\eta}\right)^{N\bar{x}} \left(1 - \frac{\eta}{1+\eta}\right)^{N(1-\bar{x})}$$

$$\hat{\eta} = \underset{\eta}{\operatorname{argmax}} L(\eta)$$

$$\frac{1}{1+\eta}$$

① Get $\frac{\partial \ell}{\partial \eta}$.

$$\ell(\eta) = \log L(\eta) = N\bar{x} \log\left(\frac{\eta}{1+\eta}\right)$$

$$+ N(1-\bar{x}) \log\left(\frac{1}{1+\eta}\right)$$

$$= N\bar{x} [\log(\eta) - \log(1+\eta)]$$

$$= N\bar{x} [\log(\eta) - \log(1+\eta)] - N(1-\bar{x}) \log(1+\eta)$$

$$\frac{\partial \ell}{\partial \eta} = N\bar{x} \left[\frac{1}{\eta} - \frac{1}{1+\eta} \right] - N(1-\bar{x}) \left(\frac{1}{1+\eta} \right)$$

$$= \frac{N\bar{x}}{\eta} - \frac{N\bar{x}}{1+\eta} - \frac{N}{1+\eta} + \frac{N\bar{x}}{1+\eta}$$

$$= \frac{N\bar{x}}{\eta} - \frac{N}{1+\eta}$$

② set $\frac{\partial \ell}{\partial \eta} = 0$ solve for η .

$$\frac{N\bar{x}}{\eta} - \frac{N}{1+\eta} = 0$$

$$\Rightarrow \frac{\bar{x}}{\eta} = \frac{1}{1+\eta}$$

$$\Rightarrow \bar{x}(1+\eta) = \eta$$

$$\Rightarrow \bar{x} + \bar{x}\eta = \eta$$

$$\Rightarrow \bar{x} = \eta(1-\bar{x})$$

$$\Rightarrow \boxed{\hat{\eta} = \frac{\bar{x}}{1-\bar{x}}} = \frac{\hat{p}}{1-\hat{p}}$$

Recall: $\eta = \frac{p}{1-p}$.

Theorem: Transf. of MLEs

If $\hat{\theta}$ is the MLE for θ then the MLE for $g(\theta)$ is $g(\hat{\theta})$.

Ex. let $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$, $\lambda > 0$.

Let's get MLE for λ .

$$\textcircled{1} L(\lambda) = f_{\lambda}(x) = \prod_{n=1}^N f(x_n)$$

ignore

$$L(\lambda) = \prod_{n=1}^N \frac{\lambda^{x_n - 1} e^{-\lambda}}{x_n!} \mathbb{1}(x_n \in \mathbb{N}_0)$$

ignore
no λ

$$= \frac{\lambda^{\sum_n x_n} e^{-N\lambda}}{\prod_n x_n!}$$

$$l(\lambda) = \left(\sum_n x_n \right) \log(\lambda) - N\lambda - \log\left(\prod_n x_n! \right)$$

$$\textcircled{2} \quad \frac{\partial l}{\partial \lambda} = 0$$

$$\frac{\partial l}{\partial \lambda} = \left(\sum_n x_n \right) \frac{1}{\lambda} - N - 0 = 0$$

$$\Rightarrow \sum_n x_n / \lambda - N = 0$$

$$\Rightarrow \hat{\lambda} = \frac{1}{N} \sum x_i = \bar{x}$$

$$\Rightarrow \left[\hat{x} = \frac{1}{N} \sum_n x_n = \bar{x} \right]$$