

Ridge Regression

Ridge penalizes the sq. err. to avoid "large" β :

$$\hat{\beta}^{(\text{ridge})} = \underset{\beta}{\operatorname{argmin}} \underbrace{\|Y - X\beta\|_2^2}_{\text{sq. err.}} + \underbrace{\lambda \|\beta\|_2^2}_{\text{penalty}}$$

$\lambda \geq 0$

By adding $\lambda \|\beta\|_2^2$ if the entries of β become large so does this penalty

λ = penalty strength

$\lambda = 0$ gives OLS $\hat{\beta}$

$\lambda \rightarrow \infty$ we get $\hat{\beta}^{(\text{ridge})} \rightarrow 0$

Typically, we don't include β_0 in the

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Often, we standardize vars before ridge

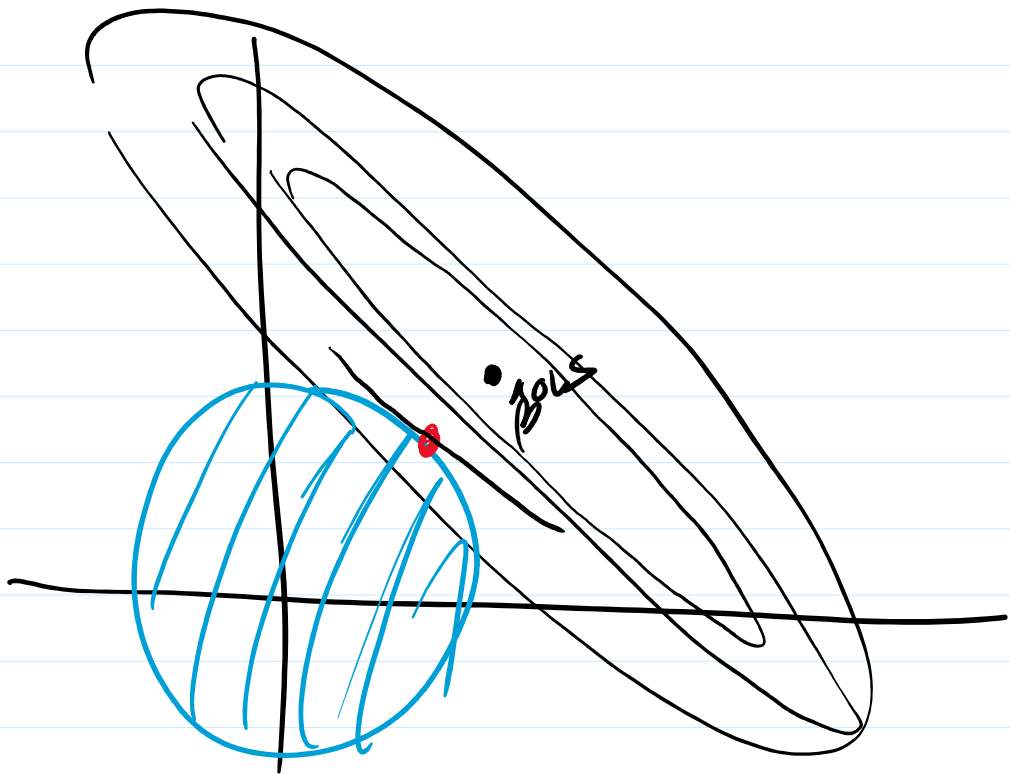
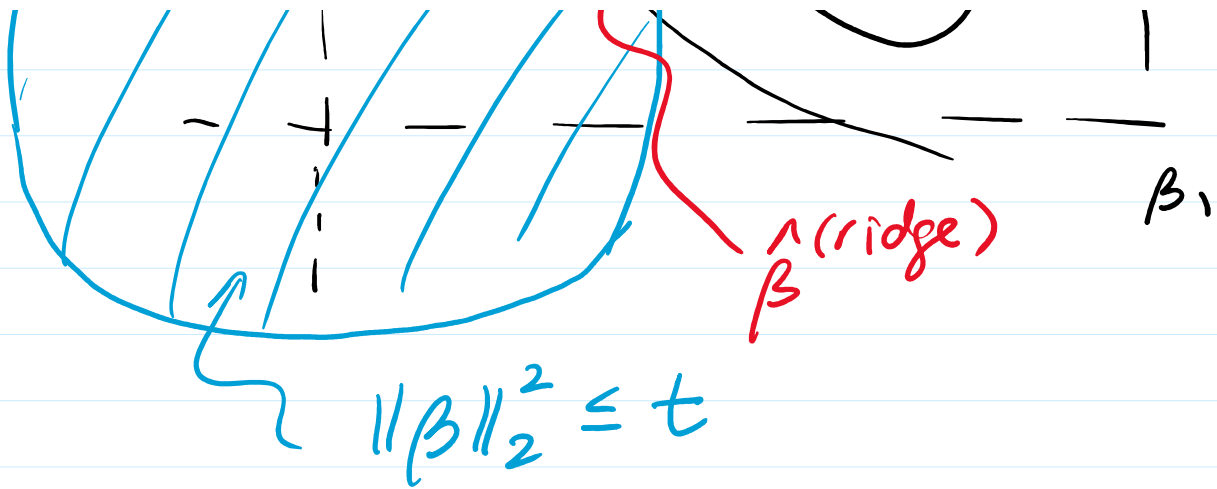
Also, typically choose λ via χ -val.

Second interpretation: ridge is equivalent to

$$\hat{\beta}(\text{ridge}) = \underset{\beta}{\operatorname{argmin}} \|Y - X\beta\|_2^2$$

$$\text{s.t. } \|\beta\|_2^2 \leq t$$





How do we get $\hat{\beta}(\text{ridge})$?

Because $\|\beta\|_2^2$ is quadratic and so is

$\frac{P}{D} \cdot 2$

$$\rightarrow \sum_{j=1}^p \beta_j^2$$

$\|Y - X\beta\|_2^2$ then there is a closed form
Solu for $\hat{\beta}^{(ridge)}$.

OLS: $\frac{\partial L}{\partial \beta} = 0 \Rightarrow \text{Solve } (X^T X) \beta = X^T Y$

Ridge: $\frac{\partial L}{\partial \beta} = 0 \Rightarrow \text{Solve } (X^T X + \lambda I) \beta = X^T Y$

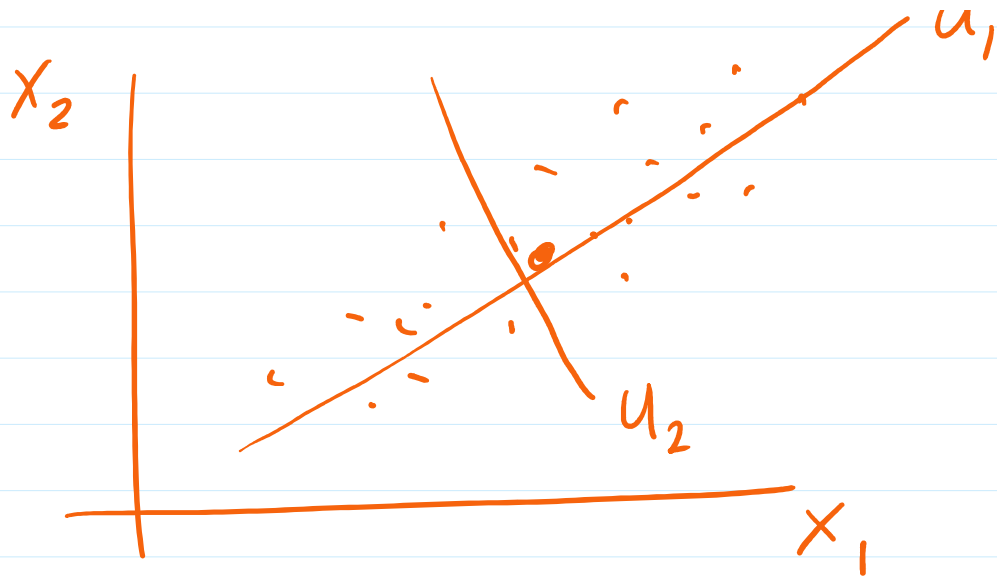
for $\lambda > 0$ $X^T X + \lambda I$ is invertible

So $\hat{\beta}^{(ridge)} = (X^T X + \lambda I)^{-1} X^T Y$.

For OLS the sens. of $\hat{\beta}^{(OLS)}$ depended on
 $K(X^T X)$

For ridge the sens. of $\hat{\beta}^{(ridge)}$ depr.

... (... T ...)



For ridge : can show that

$$\hat{Y} = X\hat{\beta}^{(ridge)} = \sum_{j=1}^p \left(\frac{\sigma_j^2}{\sigma_j^2 + \lambda} \right) u_j u_j^T Y$$

- ① project Y onto u_j
 - ② rescale each by $\frac{\sigma_j^2}{\sigma_j^2 + \lambda} \leq 1$
 - ③ sum up contribs
- \nearrow
 scale comps
 assoc. w/ smaller
 σ_j^2 : more towards

σ_j more towards zero

Degrees of Freedom

For OLS : $df = P$ if $\text{rank}(X) = P$

For ridge: $df = \sum_{j=1}^P \frac{\sigma_j^2}{\sigma_j^2 + \lambda} \leq P$

$df \rightarrow 0$ as $\lambda \rightarrow \infty$

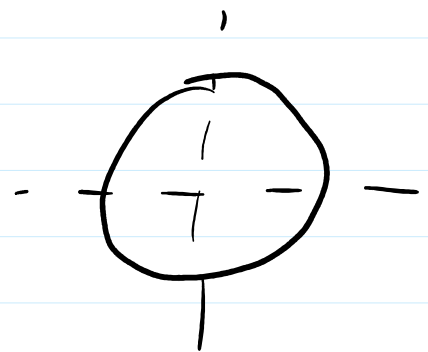
$df \rightarrow P$ as $\lambda \rightarrow 0$

Norms

Euclidean Norm:

$$\|x\|_2 = \sqrt{\sum_{j=1}^P x_j^2}$$

Consider: $\{x : \|x\|_2 = 1\}$



Can generalize:

$$q\text{-norm} : \|x\|_q = \left(\sum_{j=1}^P |x_j|^q \right)^{1/q}$$

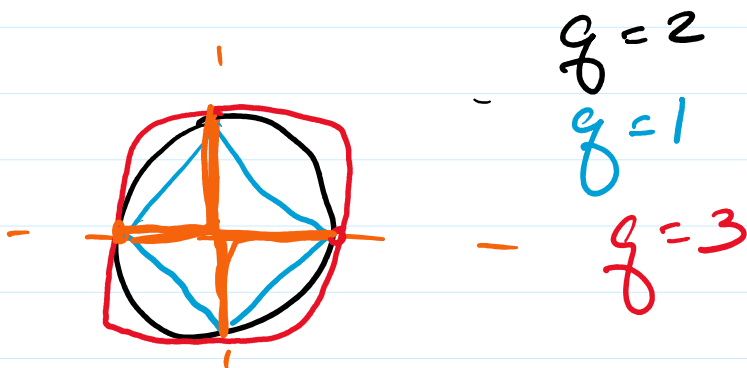
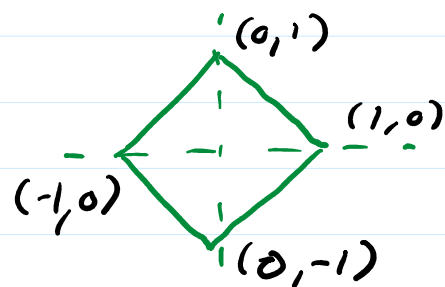
$$q\text{-norm: } \|x\|_q = \left(\sum_{j=1}^p |x_j|^q \right)^{1/q}$$

When $q=2$, I get euclidean norm
(L2 norm)

If $q=1$, get L1 norm

$$\|x\|_1 = \sum_{j=1}^p |x_j|$$

Consider $\{x : \|x\|_1 = 1\}$



as $q \rightarrow \infty$ I get $\|x\|_q \rightarrow \max_j |x_j| = \|x\|_\infty$

$q \rightarrow 0$ I get $\|x\|_q \rightarrow \#$ of non-zero

$q \rightarrow 0$ I get $\|X\|_q \rightarrow \#$ of non-zero elements
 $= \|X\|_0$

Variable selection is like zeroing out some of my $\hat{\beta}_s$:

$$\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \dots$$

\uparrow set $\hat{\beta}_2 = 0$

now: $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots$
